## Ernst Lykke Jensen Holger Rootzén

## A Note on de Moivre's

## Limit Theorems: Easy Proofs



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EASY PROOFS

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A note on de Moivre's limit theorems: Easy proofs.
by
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We consider the standardized binomial distribution

$$
f_{n}\left(x_{n, j}\right)=\sqrt{n p q}(\underset{j}{n}) p^{j} q^{n-j},
$$

where

$$
x_{n, j}=\frac{j-n p}{\sqrt{n p q}}, j=0,1, \ldots, n ; n=1,2, \ldots
$$

and $p$ is fixed, $0<p<1, p+q=1$. Together with the points $\left(x_{n, j}, f_{n}\left(x_{n, j}\right)\right)$, $j=-1,0,1, \ldots, n, n+1$, we take the lines connecting these points and continue to call this broken line $f_{n}$. Now let $n$ and $j$ tend to infinity in such a way that $x_{n, j}$ tends to a number $x$, say. Then the sequence $f_{n}$ has the Gaussian density $f(x)=\exp \left(-\frac{1}{2} x^{2}\right) / \sqrt{2 \pi}$ as 1imit function. This is, of course, wel1known; in fact, it is de Moivre's celebrated 1imit theorem. The proof is usual1y based on Stirling's formula or Fourier transformation. In what follows we give a proof based on the sequence $f_{n}^{\prime}$ of derivatives. We think that a rigorous proof along these lines ought to be available in the literature, but have

[^0]only come across heuristic arguments: hence the present note.

We begin the proof by observing that it is possible to choose the sequence $n$ and corresponding $j$ such that for large $n$ the given $x$ is located between $x_{n, j}$ and $x_{n, j+1}$. Then the derivative $f_{n}^{\prime}(x)$ is equal to the slope of the line segment connecting $\left(x_{n, j}, f_{n}\left(x_{n, j}\right)\right.$ and $\left(x_{n, j+1}, f_{n}\left(x_{n, j+1}\right)\right.$. The relative slope of this line segment is then

$$
\frac{f_{n}\left(x_{n, j+1}\right)-f_{n}\left(x_{n, j}\right)}{\left(x_{n, j+1}-x_{n, j}\right) f_{n}\left(x_{n, j}\right)}=-\frac{n p x_{n, j}+\sqrt{n p q}}{n p+\sqrt{n p q} x_{n, j}+1},
$$

which tends to $-x$, when $n$ and $j$ tends to infinity in the manner described above, and the convergence is uniform on bounded intervals. [A heuristic proof might stop here by noting that $-x$ is also the relative derivative of the Gaussian density at $x$ ]. Hence it is enough to show that

$$
\lim _{n \rightarrow \infty} \frac{f_{n}^{\prime}(x)}{f_{n}(x)}=-x
$$

uniformly on bounded intervals, implies $f_{n}$ converges to the Gaussian density.

Since the convergence is uniform, the operations of taking limit and integration can be interchanged, i.e.

$$
\begin{aligned}
\lim _{n \rightarrow \infty} \int_{0}^{u} \frac{f_{n}^{\prime}(x)}{f_{n}(x)} d x & =\int_{0}^{u} \lim _{n \rightarrow \infty} \frac{f_{n}^{\prime}(x)}{f_{n}(x)} d x \\
& =\int_{0}^{u}(-x) d x \\
& =-\frac{1}{2} u^{2} .
\end{aligned}
$$

On the other hand, $f_{n}$ is continuous and $f_{n}^{\prime}$ is a step function, so that $f_{n}$ and then $\log f_{n}$ can be recovered from their derivatives by integration. Hence

$$
\begin{aligned}
\lim _{n \rightarrow \infty} \int_{0}^{u} \frac{f_{n}^{\prime}(x)}{f_{n}(x)} d x & =\lim _{n \rightarrow \infty} \int_{0}^{u} d \log f_{n}(x) \\
& =\lim _{n \rightarrow \infty} \log \frac{f_{n}(u)}{f_{n}(0)}
\end{aligned}
$$

and since log is continuous, we have

$$
\begin{aligned}
\log \lim _{n \rightarrow \infty} \frac{f_{n}(u)}{f_{n}(0)} & =\lim \log \frac{f_{n}(u)}{f_{n}(0)} \\
& =-\frac{1}{2} u^{2}
\end{aligned}
$$

i.e.
(1)

$$
\lim _{n \rightarrow \infty} \frac{f_{n}(u)}{f_{n}(0)}=e^{-\frac{1}{2} u^{2}}
$$

Since the convergence in (1) is uniform on bounded intervals, it follows that
(2)

$$
\begin{aligned}
\lim _{n \rightarrow \infty} \int_{a}^{b} \frac{f_{n}(u)}{f_{n}(0)} d u & =\int_{a}^{b} \lim _{n \rightarrow \infty} \frac{f_{n}(u)}{f_{n}(0)} d u \\
& =\int_{a}^{b} e^{-\frac{1}{2} u^{2}} d u
\end{aligned}
$$

In particular,
(3)

$$
\begin{aligned}
\lim _{a \rightarrow \infty} \lim _{n \rightarrow \infty} \int_{-a}^{a} \frac{f_{n}(u)}{f_{n}(0)} d u & =\lim _{a \rightarrow \infty} \int_{-a}^{a} e^{-\frac{1}{2} u^{2}} d u \\
& =\sqrt{2 \pi} .
\end{aligned}
$$

Since $f_{n}$ is a probability density for a random variable with expectation and variance tending to the expectation and variance, 0 and 1 respectively, of the standardized binomial distribution, an application of Tchebycheff's inequality yields

$$
1-\frac{1}{a^{2}} \leqq \lim _{n \rightarrow \infty} \inf \int_{-a}^{a} f_{n}(u) d u \leqq \lim _{n \rightarrow \infty} \sup _{-a}^{a} f_{n}(u) d u \leqq 1
$$

Hence from (3)
(4) $\quad \lim _{n \rightarrow \infty} \frac{1}{f_{n}(0)}=\sqrt{2 \pi}$.
and the assertion now follows from (1).

In a short-hand notation, we have proved de Moivre's first result

$$
\binom{n}{j} p^{j} q^{n-j} \sim \frac{1}{\sqrt{2 \pi n p q}} e^{-\frac{(j-n p)^{2}}{2 n p q}}
$$

where the symbol $\sim$ means that the ratio between the 1 eft-hand side and the right-hand side tends to 1.

It is easily seen that

$$
\sum_{j: a \leqq x_{n, j} \leqq b} \frac{f_{n}\left(x_{n, j}\right)}{\sqrt{n p q}}
$$

differs from

$$
\int_{a}^{b} f_{n}(x) d x
$$

by at most $\max _{j} f_{n}\left(x_{n, j}\right) / \sqrt{n p q}$. Hence, using (2) and (4), we have de Moivre's second result,

$$
\sum_{j=a}^{b^{\prime}}\binom{n}{j} p^{j} q^{n-j} \sim \int_{a}^{b} \frac{1}{\sqrt{2 \pi}} e^{-\frac{1}{2} x^{2}} d x
$$

for

$$
\frac{a^{\prime}-n p}{\sqrt{n p q}} \rightarrow a, \frac{b^{\prime}-n p}{\sqrt{n p q}} \rightarrow b
$$

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