

Monday December 6 2004

We recall the following important result

Theorem Suppose P is irreducible, aperiodic and Harris recurrent. Then there is a small set $C \in \mathcal{E}^+$ such that

$$(1) \quad \sup_{x \in C} H_{C,x}(X) < \infty$$

then P has an invariant probability measure.

□

The left hand side in (1) has a probabilistic interpretation: if X_0, X_1, \dots is a P -chain with $X_0 = x$, and if τ_C is the first return time to C , then

$$E(\tau_C) = H_{C,x}(X)$$

hence the condition is frequently written as

$$\sup_{x \in C} E_x(\tau_C) < \infty$$

where the symbol E_x means the expected value if the initial distribution is δ_x .

We will make this result operational through a drift criterion. The key result in this process is a triviality known as Dobinski's Lemma

Dobinski's Lemma: Let $\mathbb{F}_0, \mathbb{F}_1, \dots$ be a filtration

Let $(Y_n)_{n \geq 0}$, $(Z_n)_{n \geq 0}$ and $(W_n)_{n \geq 0}$ be the non-negative processes. Suppose the Z -process is integrable and that

$$Y_n \leq Z_n - \mathbb{E}(Z_{n+1} | \mathbb{F}_n) + W_n \quad \forall n$$

Then

$$\mathbb{E} \sum_0^{\bar{\sigma}} Y_n \leq \mathbb{E} Z_0 + \mathbb{E} \sum_0^{\bar{\sigma}} W_n$$

for any stopping time $\bar{\sigma}$

Proof: By summation, we obtain

$$\sum_0^{\bar{\sigma}} Y_n \leq \sum_0^{\bar{\sigma}} Z_n - \mathbb{E}(Z_{n+1} | \mathbb{F}_n) + \sum_0^{\bar{\sigma}} W_n$$

$$= Z_0 + \sum_1^{\bar{\sigma}} Z_n - \mathbb{E}(Z_n | \mathbb{F}_{n-1}) - \mathbb{E}(Z_{\bar{\sigma}+1} | \mathbb{F}_{\bar{\sigma}}) + \sum_0^{\bar{\sigma}} W_n$$

Introduce the zero-mean martingale

$$M_n = \sum_1^n Z_k - \mathbb{E}(Z_k | \mathcal{F}_{k-1})$$

Since that $Z_{n+1} \geq 0$, and thus $\mathbb{E}(Z_{n+1} | \mathcal{F}_n) \geq 0$ a.s.,
we obtain that

$$\sum_0^n Y_k \leq Z_0 + M_n + \sum_0^n W_k \quad \text{a.s.}$$

and thus

$$\mathbb{E} \sum_0^\delta Y_k \leq \mathbb{E} Z_0 + \mathbb{E} M_\delta + \mathbb{E} \sum_0^\delta W_k$$

If we only consider a bounded stopping time δ
The theorem on optional sampling shows that M_δ
has mean zero. And thus

$$\mathbb{E} \sum_0^\delta Y_k \leq \mathbb{E} Z_0 + \mathbb{E} \sum_0^\delta W_k$$

If δ is a semi-stopping time - an infinite -
we have that

$$\mathbb{E} \sum_0^{\text{D.W.}} Y_k \leq \mathbb{E} Z_0 + \mathbb{E} \sum_0^{\text{D.W.}} W_k$$

for every N . As

$$E \sum_0^{2N} Y_n \rightarrow E \left(\sum_0^{\infty} Y_n \right)$$

$$E \sum_0^{2N} Z_n \rightarrow E \left(\sum_0^{\infty} Z_n \right)$$

by dominated convergence, the result follows

□

Theorem: Let P be irreducible and aperiodic. Suppose there is a drift function $V: X \rightarrow [0, \infty)$ such that

$$(b) \quad \Delta V(x) \leq -1 + b \mathbb{1}_C(x)$$

for some small set C on which V is bounded. Then we conclude

1) P is Harris

2) P possesses an invariant probability measure

Proof: Write out (b) since that

$$PV(x) \leq V(x) - 1 \quad x \notin C$$

$$PV(x) \leq V(x) + b - 1 \quad x \in C$$

At test

$$PV(x) \leq V(x) + M \quad \forall x$$

for some M . This implies that

$$P^k V(x) \leq V(x) + kM$$

The core of the market is that if x_0, x_1, \dots is a path with $\mathbb{E} V(x_0) < \infty$, then

$$\begin{aligned} \int V(x_n) dP &= \int P^k V(x_0) dP \\ &\leq \int V(x_0) + kM dP \\ &< \infty \end{aligned}$$

Note that (a) says

$$1 \leq V(x) - PV(x) + b/c(x) \quad \forall x$$

a

$$1 \leq V(x_n) - PV(x_n) + b/c(x_n) \quad \forall n$$

Note that $PV(x_n) = \mathbb{E}(V(x_{n+1}) | x_0, \dots, x_n)$, so

Dobinski's lemma implies that

$$(A) \quad \mathbb{E} \sum_{i=0}^{\infty} 1/c(x_i) \leq V(x_0) + b \sigma \sum_{i=0}^{\infty} 1/c(x_i)$$

This holds for any stopping time, but we will for now use it with $T = T_C$ because that diagonalizes the left sum. We obtain

$$\bar{B}_x T \leq V(x) + Z(x) - 1$$

As a first application, $\bar{B}_x T < \infty$ which certainly implies $L_x(C) \neq \emptyset$ for all x . We conclude from this that C is Harris recurrent and $\mathbb{E}T$ and this is quite sufficient to conclude that P is Harris recurrent.

Next we can try to

$$\sup_{x \in C} \bar{B}_x T \leq \sup_{x \in C} V(x) + Z(x) - 1 < \infty$$

and thus P is positive

□

This drift criterion is very, very useful because it is frequently easy to check. It gives the existence of an invariant probability measure, but it does not really give a clue to what this measure is.

Example: Consider an autoregressive scheme on \mathbb{R}

$$X_{n+1} = \rho X_n + U_{n+1}$$

wh $|\rho| < 1$ and wh U_1, U_2, \dots are iid, we insist

$$\mathbb{E}(U_i) = 0 \quad \mathbb{E} U_i^2 < \infty$$

had we insist irreducibility and aperiodicity has been checked and that compact sets are small.

Consider the drift function $V(x) = cx^2$.

$$\begin{aligned} P V(x) &= \mathbb{E}(V(\rho x + U)) = c \mathbb{E}(\rho^2 x^2 + 2\rho x U + U^2) \\ &= c(\rho^2 x^2 + 2\rho x \mathbb{E}U + \mathbb{E}U^2) \\ &= \rho^2 c x^2 + c \mathbb{E}U^2 \end{aligned}$$

Find some λ such that $\rho^2 < \lambda < 1$.

$$\rho^2 V(x) + c \mathbb{E}U^2 \leq \lambda V(x)$$

$$\Leftrightarrow c \mathbb{E}U^2 \leq (\lambda - \rho^2) V(x)$$

$$\Leftrightarrow \frac{c \mathbb{E}U^2}{\lambda - \rho^2} \leq V(x)$$

$$\Leftrightarrow \sqrt{\frac{c \mathbb{E}U^2}{\lambda - \rho^2}} \leq |x|$$

Similarly,

$$\lambda v(x) \geq v(x) - 1$$

$$1 \leq (1-\lambda)v(x)$$

$$\sqrt{\frac{1}{c(1-\lambda)}} \leq |x|$$

hence, for $|x| \geq c = \max \left\{ \sqrt{\frac{b\sigma^2}{\lambda - \rho}}, \sqrt{\frac{1}{c(1-\lambda)}} \right\}$

we have that

$$Pv(x) \leq v(x) - 1 \quad x \notin [L-c, c]$$

And clearly, both v and Pv are bounded on $[L-c, c]$, so

$$Pv(x) \leq v(x) + b \quad x \in [L-c, c]$$

for some suitable b

So the AR(1) process satisfies the drift condition for positive values.

Alternativ kann man direkte 2. moment schätzen, es sei

$$p: \quad V(x) = |x|. \quad \text{Da}$$

$$P V(x) = \sigma |p x + 0| \leq |p| |x| + \sigma |U|$$

es gilt also

$$|p| |x| + \sigma |U| \leq |x| - 1 \quad \Rightarrow$$

$$1 + \sigma |U| \leq (1 - |p|) |x| \quad \Rightarrow$$

$$\frac{1 + \sigma |U|}{1 - |p|} \leq |x|$$

Es ist also

$$P V(x) \leq V(x) - 1 \quad p \in [-c, c]^c$$

$$\text{Nun } c = \frac{1 + \sigma |U|}{1 - |p|} \quad \text{Reversiert } p \in [-c, c]$$

zu zeigen.

Second example: Markov process,

$$X_{n+1} \approx \sqrt{a^2 + b^2 X_n^2} \cdot U_{n+1}$$

ergodic if $b^2 < 1$, $\sigma |U| < \infty$

Theorem: If P satisfies the criteria of the previous theorem, then C is regular, that is

$$\sup_{x \in C} \beta_x \tau_B < \infty \quad \forall B \in \mathbb{E}^+$$

Proof: Essentially this follows from formula (10) which reads that

$$\beta_x \tau_B \leq V(x) - 1 + b \beta_x \sum_0^{\tau_B} |c(x_n)|$$

Since $C \in \mathbb{E}^+$ is small, there is $k, \delta > 0$ such that

$$r_x^k(B) \geq \delta \quad \forall x \in C$$

Each time the process visits C there is a lottery involved, i.e. B will be visited k steps later. If the success probability is taken to be δ , these lotteries will be independent and hence the expected number of lotteries before success will be $\frac{1}{\delta}$.

Hence

$$\beta_x \sum_0^{\tau_B} |c(x_n)| \leq \frac{1}{\delta} \sup_{x \in C} \beta_x \tau_C + k$$

for all $x \in C$.

(proof not entirely accurate).

□

Theorem if P satisfies the criterion of the previous theorem, the sublevel sets of V are all small.

proof: let $C_n = \{x \mid V(x) \leq n\}$. Using (1) we see that for $x \in C_n$

$$\begin{aligned} \delta_x \tau_c &\leq V(x) + b \delta_x \sum_0^{\tau_c} 1_{(x, \tau_c)} - 1 \\ &\leq n + 2b - 1 \end{aligned}$$

Hence

$$P_x(\tau_c \geq n) \leq \frac{\delta_x \tau_c}{n} \leq \frac{n + 2b - 1}{n} < \frac{1}{2}$$

if n is large enough, irrespective of x . Thus

$$P_x^k(c) > \frac{1}{2m} \quad \text{for at least one } k=1, \dots, m \quad \forall x \in C_n$$

This implies that C_n $\overset{a}{\text{meets}} C$ when the sampling distribution is the uniform distribution on $1, \dots, m$. Hence it follows that C_n is petite. And periodically implies that C_n is small.

□

corollary: $X = \cup C_n$, each C_n recurrent.

□